The new benchmark for forecasts of the real price of crude oil

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What is a good forecast?

Suppose we want to evaluate a forecast of X_{t+h}

Common practice: compare accuracy of your forecast against

no-change benchmark $\equiv X_t$

- ► X_t is optimal forecast under the random walk hypothesis → accuracy-improvements over NCB imply that
 - o series is predictable in general
 - $\circ~$ our forecast is more useful than "naive" approach

But: this approach not informative for temporally aggregated data

 lower-frequency series constructed by *averaging* or summation (e.g., real commodity prices, interest rates, ...)

The new benchmark

Random walk hypothesis applies to high-frequency series

- aggregated data are predictable by construction (Working 1960)
- $\circ\,$ conventional NCB is not the optimal forecast
- \rightarrow improvements over conventional NCB not informative about RWH
- New no-change benchmark:

last high-frequency observation

- $\circ\,$ optimal under random walk hypothesis
 - \rightarrow restores original interpretation of comparison with NCB
- $\circ~\approx$ 45% improvement in MSPE for monthly/quarterly averages of daily data for 1-step-ahead prediction

Application to the real price of crude oil

- Real price of crude oil typically based on averaged data (Kilian 2009; Baumeister & Hamilton 2019)
- Existing literature: model-based forecasts beat NCB (Baumeister & Kilian 2012, 2014, 2015; Alquist et al. 2013; Snudden, 2018; Funk, 2018; Garratt et al., 2019)
- New no-change benchmark based on monthly closing price of oil

 ≈ 40% improvements in accuracy at 1-step-ahead prediction
 using closing prices for estimation improves traditional models
 but: most models do *not* beat the new benchmark
 - \rightarrow oil prices are more difficult to predict than previously thought

Intuition under the RW null hypothesis

Forecaster's goal is to predict X_{t+h} given time t information

$$X_t \equiv \frac{1}{n} \sum_{i=1}^n y_{t,i}$$

t =month; i =day of month; n = # of days in month

Null hypothesis: daily observations follow random walk

$$y_{t,i} = y_{t,i-1} + \epsilon_{t,i}, \quad \text{ for } i = 1, \cdots, n$$

 $\epsilon_{t,i}$ is a mean-zero, *iid* error term with variance σ_{ϵ}^2

- What is the optimal forecast in this setting?
- What are the consequences of using X_t to evaluate forecasts?

The optimal forecast under the RWH

Optimal forecast in MSPE terms is the conditional expectation

$$E_t(X_{t+h}) = E(X_{t+h}|y_{1,1}; \cdots; y_{t,n}; Z_1; \cdots; Z_t)$$

RWH: conditional expectation of each future daily obs

$$E_t(y_{t+h,i})=y_{t,n}$$
 for all $h>0,\;i=1,\cdots,n$

 \Rightarrow conditional expectation of each future average obs

$$E_t(X_{t+h}) = E_t(\frac{1}{n}\sum_{1}^{n}y_{t+h,n}) = y_{t,n}$$

New benchmark is the optimal forecast under the RWH:

last high-frequency observation in period t
$$\equiv y_{t,r}$$

Comparing the no-change forecasts X_t and $y_{t,n}$

▶ What is the MSPE of y_{t,n} relative to the conventional NCB X_t?

MSPE for conventional benchmark:

$$E\left[(X_t - X_{t+h})^2\right] = \left((h-1) \cdot n + \frac{(n+1) \cdot (2n+1)}{6n} + \frac{(n-1) \cdot (2n-1)}{6n}\right)\sigma_{\epsilon}^2$$

MSPE for new benchmark:

$$E\left[\left(y_{t,n}-X_{t+h}\right)^{2}\right]=\left((h-1)\cdot n+\frac{(n+1)\cdot(2n+1)}{6n}\right)\sigma_{\epsilon}^{2}$$

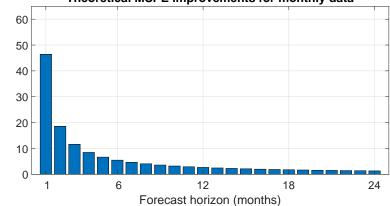
▶ y_{t,n} has lower MSPE than X_t under the RWH

All forecasting gains occur at the 1-step ahead prediction

• forecast gain is constant, forecast error increases with $h \Rightarrow$ relative improvements largest for short horizons

MSPE ratios of the two NCBs

% MSPE improvement



Theoretical MSPE improvements for monthly data

Consequences for forecast comparisons

- Conventional NCB is not the optimal forecast under the RWH
- Moreover, ΔX_t is autocorrelated (Working 1960)

$$X_t - X_{t-1} = \frac{1}{n} \left[\sum_{i=1}^n i \cdot \epsilon_{t-1,i} + \sum_{j=1}^n (n+1-j) \cdot \epsilon_{t,j} \right]$$

 \Rightarrow improvements over no-change forecast X_t are expected *even under* the random walk null hypothesis

 \Rightarrow improvements over no-change forecast $y_{t,n}$ are evidence against *the* random walk null hypothesis

- Using $y_{t,n}$ as a benchmark
 - maintains original spirit of comparisons with the NCB
 - $\circ\,$ more difficult to achieve when HF observations are persistent

Application: Forecasting the real price of crude oil

- Forecast of the monthly real price of crude oil in standard setting (Baumeister & Kilian 2012, 2014, 2015; Alquist et al. 2013)
 - $\circ~$ real oil price is deflated average of daily nominal prices

$$\bar{p}_t^r = \frac{\frac{1}{n} \sum_{i=1}^n p_{t,i}}{CPI_t}$$

 $\circ\,$ goal: forecast \bar{p}_{t+h}^r given month t information

The new benchmark: series of monthly closing prices

$$p_t^{r,closing} = rac{p_{t,n}}{CPI_t}$$

Econometric models

- 1. Univariate time series models
 - \circ AR (log-level and percent changes)
 - ARMA(1,1)
 - ARFI
- 2. VAR models (Kilian & Murphy 2014)
 - $\circ~$ percent change in global crude oil production
 - real economic activity indicator (Kilian 2009)
 - $\circ~$ real price of oil
 - $\circ~$ change in above-ground global crude oil inventories
 - > unrestricted least-squares estimation
 - > Gaussian BVAR with prior variance (Giannone, Lenza & Primiceri 2010)
- 3. Futures price curve
- 4. Equal-weight forecast combination

Implementation

Following Baumeister & Kilian (2012):

- Real-time data
- Out-of-sample forecasts with expanding window
 - estimation period starts 1973M2
 - $\circ~$ out-of-sample evaluation 1992M1 2018M12
- All models estimated with average prices

Criteria: MSPE ratio and directional accuracy

- displayed relative to conventional NCB
- $\circ\,$ improvements over the new NCB are in \boldsymbol{bold}
- standard tests for inference

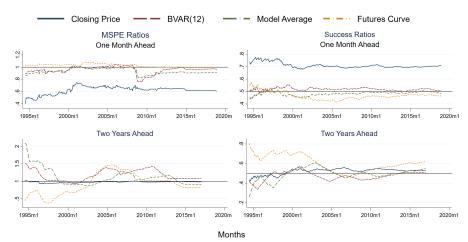
(Diebold & Mariano 1995, Pesaran and Timmermann 2009)

Baseline results: Real WTI prices, real-time data

Months Ahead	Last WTI Close Price	BVAR(12)	VAR(12)	AR(12)	AR(12) %∆	ARFI(1)	ARMA(1,1)	Futures Curve	Model Averaging
MSPE Ratios									
1	0.61***	0.97	1.01	0.94	0.95*	0.93	0.92	1.00	0.91
3	0.89**	1.00	1.00	0.97	0.99	0.96	0.95	0.97	0.92
6	0.95**	1.05	1.04	1.00	1.04	0.99	0.95	0.96	0.95
12	0.96**	1.10	1.11	1.00	1.10	1.00	0.94	0.85**	0.94
24	0.99	1.08	1.06	1.02	1.19	1.04	0.96	0.82*	0.91
Success Ratios									
1	0.71***	0.51	0.54**	0.52	0.49	0.52	0.53	0.47	0.49
3	0.60***	0.53	0.54*	0.49	0.57**	0.49	0.50	0.49	0.52
6	0.56**	0.53	0.55*	0.48	0.54	0.49	0.46	0.53	0.51
12	0.59***	0.49	0.56**	0.53	0.50	0.51	0.50	0.61***	0.51
24	0.53	0.51	0.55	0.57	0.47	0.55	0.56	0.62***	0.55

Note: ***, **, and * denote significant improvement over the average no-change forecast at the 1%, 5%, and 10% level. Bold values indicate significant improvements over the last closing price no-change forecast at the 5 percent level.

Evolution of baseline WTI Real price forecasts: Real-time data



Note: Dynamic, recursive, out-of-sample forecasts 1992M1–2018M12. The forecast criteria reported include the recursive MSPE expressed as a ratio relative to the monthly average no-change forecast. All forecast criteria are evaluated in the levels of the real price of oil. The first 30 months are dropped to reduce starting-point effects.

Interpreting baseline results

- Large(!) forecasting gains for short forecast horizons
- Actual vs. predicted MSPE ratios (RWH for daily data)

Horizon (months)	1	3	6	12	24
Theoretical	0.54	0.88	0.95	0.97	0.99
Empirical (revised data)	0.60	0.89	0.95	0.96	0.99

Robustness: remarkably similar results for

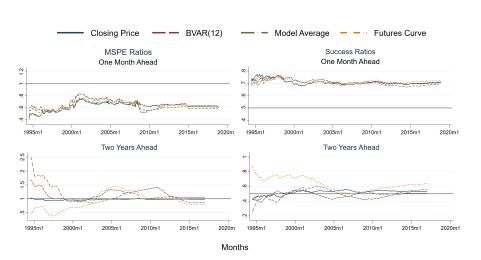
- $\circ\,$ various estimation / evaluation periods
- o different oil price series (Brent, RAC, nominal prices)
- $\circ~$ ex-post revised data, nominal price, quarterly data
- What about estimating models with closing prices?

Models estimated with last closing price

Months	Last WTI	BVAR(12)	VAR(12)	AD(2)	AR(1)	ADEI(2)	ADMA(1.1)	Last Future	Model
Ahead	Close Price	BVAR(12)	VAR(12)	AR(2)	%Δ	ARFI(2)	ARMA(1,1)	Price Curve	Averaging
	MSPE Ratios								
1	0.61***	0.62**	0.73*	0.57***	0.62***	0.57***	0.57***	0.60***	0.58***
3	0.89**	0.90	0.91	0.85	0.93*	0.86	0.86	0.88**	0.82*
6	0.95**	0.98	0.97	0.89	1.04	0.89	0.90	0.92*	0.87
12	0.96**	1.05	1.06	0.90	1.15	0.90	0.91	0.84**	0.87*
24	0.99	1.06	1.06	0.95	1.48	0.95	0.97	0.80*	0.87
	Success Ratios								
1	0.71***	0.71***	0.66***	0.71***	0.73***	0.72***	0.71***	0.73***	0.70***
3	0.60***	0.58**	0.56**	0.57**	0.62***	0.60***	0.60***	0.60***	0.59***
6	0.56**	0.58**	0.58**	0.52	0.58**	0.51	0.52	0.55	0.55
12	0.59***	0.54	0.58**	0.54	0.59**	0.51	0.53	0.63***	0.54
24	0.53	0.51	0.57*	0.58	0.56	0.55	0.54	0.64***	0.56

Note: ***, **, and * denote significant improvement over the average no-change forecast at the 1%, 5%, and 10% level. Bold values indicate significant improvements over the last closing price no-change forecast at the 5 percent level.

Evolution of real-time closing-price WTI forecasts



Note: Dynamic, recursive, out-of-sample forecasts 1992M1–2018M12. The forecast criteria reported include the recursive MSPE expressed as a ratio relative to the monthly average no-change forecast. All forecast criteria are evaluated in the levels of the real price of oil. The first 30 months are dropped to reduce starting-point effects.

Take-away

- Aggregating higher-frequency data
 - introduces loss of information (e.g., Rossana & Seater 1995)
 - changes interpretation of standard forecast comparisons
- $\rightarrow\,$ use closing observations for estimation & forecast evaluation
 - New benchmark for real price of crude oil changes assessments of models & oil price predictability:
 - o daily oil prices are "random walkish"
 - $\circ\,$ real price of crude oil is predictable by construction
 - $\circ~$ but most models do not beat the new benchmark

Averaging can be desirable, but watch out for settings in which

- series of interest is temporally aggregated
- underlying data is persistent

Thank you!